(Following the comments presented by the referees, we have added much explanation and clarification into Secs. 2.2, 3, 4.1, 4.2 and 4.3 of our manuscript. We apologize that this has taken much time and our answer is delayed. The revised sections have been attached.)

(Our response is in the following given in italics)

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Interactive comment on “A universal error source in past climate estimates derived from tree rings” by Juhani Rinne et al.

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This paper introduces a “new” standardisation method to construct chronologies. It tries to show that there is low-frequency bias in tree ring chronology reconstructions. The new reconstruction method is used with the Tornetrask MXD data from 1988 to demonstrate this bias.

The authors state that “The presented method to estimate past temperatures from tree ring measurements is a new approach, where no age dependence of the tree rings is estimated.” yet they are clearly removing the age related growth of trees as a linear trend. For each calendar year they have a few measured rings (e.g. 20 values). They effectively create further values by extrapolation including younger rings and older rings to a total of 270 (e.g. adding 250 values) with a linear age-related decay. These 270 values are then averaged together thus removing the effect of ring-width aging using the presumption of some form of liner decay of ring width with age.

The method of estimating R1 and Rn (page 4 line 17) will need a detailed explanation.

- Determination of \( r_1 \) and \( r_n \) is now explicitly described in the revised version of Sec. 2.2.

Overall it is likely to have a similar result to that of creating and fitting a linearly decaying RCS curve. In RCS the averaging and smoothing of the RCS curve tends to reduce the climate noise from the estimation of the ageing trend whereas in the proposed method using rings from a single
year which all have the same climate signal achieves this. Their conclusion that they do not remove the effect of age-related growth from their measurements is not justified as they do try to remove the age effect.

Our method is designed to estimate error variances. Especially important are the variances of the systematic errors due to the missing young and old age classes. It is natural that in their estimation the reconstruction method resembles usual conventions. However, here the impact of the missing younger and older age classes is estimated directly from the measurements and only for the average impact of the missing young and old age classes. Otherwise the trapezoidal rule is applied.

- Explanatory text on the performance of our method is included in the revised version of Sec. 4.3.

The 1988 Tornetrask MXD data were selected by age (oldest well behaved trees) from the much larger TRW data set with a view to using curve-fitting standardisation methods with sufficient replication for reconstructing medium frequency variability. An even distribution of tree rings by age in each year was not thought necessary in 1988. This is not a suitable data set to introduce or evaluate the proposed new standardisation technique.

*It is true that there were better data for such an evaluation if that were needed. From our point of view, the 1988 Torneträsk is suitable to illustrate the performance of the error formulae derived. Otherwise it is enough that our reconstruction method performs sufficiently well (the performance is illustrated with the aid of the Torneträsk data).*

- *Comparison with Briffa et al. (1992) will be reconsidered in Sec. 4.1.*

The authors need to note that Briffa et al (2009, Hughes book chapter) show that for these MXD measurements from Tornetrask the assumption of linear decay creates a bias in the reconstruction. For TRW, the assumption of linear decay would create even more bias.

*The estimates of the error variance are so high that in any case the confidence limits of our reconstruction were wide if such limits had been constructed. The linear decay is not assumed but estimated to be applicable in missing young and old age classes. It follows that there are two results of the estimation: the average impact of those age classes and the corresponding error variance.*
Confidence limits are implicitly characterized by the degrees of freedom and are discussed in the revised version of Sec. 4.3.

The Esper 2012 chronology has more trees (even after using mean tree rather than multiple cores) and less error due to sample count i.e. less noise. Is the age distribution of the Esper trees biased over time? No assessment is made of this so the presence or absence of systematic bias is not known and the comparison (and conclusions based on it) in this paper is not justified.

There is no attempt to distinguish between bias due to the age-related growth decay in tree measurements and noise created by poor replication and the authors confuse these two effects making their conclusions less valid and unhelpful.

The detection of the bias is now explained more in details, in the manuscript only a reference was given. Different Torneträsk analyses show a similar long-term variation (as far no new trees are added to the sample). Westward and eastward from Torneträsk there are no corresponding long-term oscillations in the nearby reconstructions. Such local deviations at Torneträsk (of the order of magnitude of the greenhouse warming) are climatologically impossible and therefore the Torneträsk analyses must be biased.

- The detection of the bias will be explained in Sec. 4.1 of the revised manuscript as follows:

  "In Rinne et al. 2014 it was observed that both in the nearby oceanic (August SST, Norwegian Sea, Miettinen et al. 2012) and continental (Esper et al. 2012) temperature estimates of the long term oscillations clearly and similarly differ from those derived from the Torneträsk data. Such local anomalies are climatologically impossible and therefore the mutually similar long-term oscillations in the Torneträsk reconstructions contain a bias. Accordingly, such reconstructions are suitable for our error studies.

  The differences observed are extreme being of the order of magnitude of the greenhouse warming. The long-term bias in Torneträsk reconstructions is thus detected climatologically. In our computations we estimate that climatological bias as the difference between the Torneträsk reconstruction and the corresponding Esper et al. (2012) reconstruction, the latter having a high number of trees"
- The explanation of the bias with the aid of the error source studied is given in the revised version of Sec. 4.2.

The presentation in this paper is not suitable to introduce a “new” standardisation method. A comparison of new against existing methods is needed which should include a careful assessment of errors – with separation of noise related to insufficient samples and systematic bias related to poor removal of age-dependent growth and an evaluation of error magnitude.

Our estimate of the error variance is divided in two parts which describe the error sources that you mentioned: the noise related to insufficient samples and systematic bias related to poor removal of age-dependent growth. The former part is studied in the article. The latter one is widely studied in the literature and is outside of the scope of our study.

A sample data set with sufficient samples in each year to sub-divide the data and show the effect of reducing sample counts is needed and only then can the bias due to age-trend be shown.

Our work is focused on the estimation of error variances and the error term studied depends only on the distribution of the measurements over the years and age classes. Our approach is designed to estimate the variance of that error. Otherwise it is enough that our reconstruction works sufficiently satisfactory and therefore there is no need to validate our method with other methods.

- Explanatory text on the performance of our method is included in the revised version of Sec. 4.3.

My overall assessment is that this paper requires considerable improvement before it is suitable for publication.

We wish that the clarifications and explanations made in the manuscript have make the text clearer
Dr Thomas Melvin  

**Changes to the manuscript.**

Contents

2.2 Explicit computation formulae  

The estimation of parameters $r_1$ and $r_n$ is described explicitly

3 Torneträsk case study

The original and corrected observations for 1826-38 are presented in Fig. 4. The motivation of the correction is presented in details.

4.1. Bias of the Torneträsk reconstructions

The climatological detection of the bias in the Torneträsk reconstructions is presented more specifically.  

The use of Briffa (1992) reconstruction in Fig. 3a is explained in more detail.

4.2. Explanation of the bias observed

It is pointed out that the Torneträsk case is only an illustration and application of the general error terms in Eqs. (4) and (5).  

The bias in the reconstructions is, on the basis of the description in Sec. 4.1, estimated with the aid of the reconstruction in Esper et al. (2012).

4.3 The performance of the reconstruction method applied

The performance of the reconstruction is described only to show that the error analysis is based on a sound calculation. The resulting temperature estimates are sufficiently satisfactory.  

It is explained that the reconstruction methods published in the literature do not impact the error source studied here and are thus outside of the scope of our study.  

The selection of the upper limit of the age classes is discussed.
In order to get the average taken over all age classes, the contributions of the missing young and old age classes are needed. These can be estimated as follows. First estimates of measurements are interpolated for every age class where that is possible. Then years with values in age classes of \( b=20 \) and \( b=270 \) are selected. This makes it possible to compute yearly averages over all age classes of 20 thru 270. Next yearly averages over age classes of 21 thru 270 are computed. Generally this new average is smaller but the variation between the years is strong. By computing the mean and r.m.s.e. of the yearly differences between the averages with and without \( b=20 \) we get an estimate of the eliminated age class \( b=20 \). In the next step the yearly averages over age classes of 22 thru 270 are computed and used to estimate the impact of missing age classes of \( b=21 \) and \( b=22 \) to the average of all measurements between \( b=20 \) and \( b=270 \). The computations are continued by dropping out more age classes. Similar approach is applied to estimate the impact of missing old age classes.

The estimation turned out to be more complicated if only very young age classes (\( b_1 \approx 20 \)) or most of the older age classes (\( b_n \approx 30 \)) were missing. To keep the formulae simple, the linear approximation is extended to those cases, too.

The result is that the contributions of the missing young and old age classes to the average taken over all age classes can be estimated linearly by \( M_{young} \approx r_1 (b_1 - 1) \) and \( M_{old} \approx r_n (b_n - b_{max}) \), respectively, where \( r_1 \approx 0.000128 \) and \( r_n \approx 0.000170 \). Note that \( M_{old} \) results in a negative correction. The accuracy of such corrections turned out to be low their variances being roughly \( [r_1 (b_1 - 1)]^2 \) and \( [r_n (b_n - b_{max})]^2 \). The application of the corrections are illustrated for an individual year in Fig. 2 for \( b_{max} = 300 \).

The focus here is in the long-term error. Important are the estimates of the error variances due to the missing age classes. Precise values of \( r_1 \) and \( r_n \) are not needed. The main information included in the results is the dependence of the variance terms.
on the lengths of the data gaps, \( b_{j-1} \) and \( b_{\text{max}}-b_n \). Our reconstruction method is directed to uncover that dependence explicitly in a simple way. Otherwise it is enough that the reconstruction method performs sufficiently satisfactorily. It needs not be an optimal one.

No assumption on the age function is made. The contributions of the young and old age classes can widely vary between years as indicated by their error variances. However, average corrections are needed in order to decrease the impact of systematic errors. By adding the corrections, the yearly average of the interpolated, extrapolated and measured values becomes approximately ...

3 Torneträsk case study

*(page 7, lines 1-6, the changes are given in blue)*

Klingberj and Moberg (2003) composited a series of instrumental observations made in the Tornedalen region, some 300 km from Torneträsk. Their construction begins with instrumental data from Övertorneå (1802-1838). The observation hours were not stated explicitly for 1826-38 and the authors had to assume them. The JJA temperatures are known to be sensitive to the observation hours used. If these are not precisely known, true climatic variations may remain but the average level of the temperature may become wrong. In such cases any kind of support is welcome. During 1826-38 the bias estimate in our reconstructions is rather small (\( \approx -0.25^\circ\text{C} \), Fig. 3a) and the DOF are rather high (\( \approx 11 \), Fig. 3c). Our reconstruction can therefore be applicable. If the JJA temperatures in observations are systematically increased by 1°C every year during 1826-38, the smoothed result happens to fit rather well with our reconstruction in Fig. 4. Accordingly, the corrected temperatures are supported by the reconstruction and the anomalous coldness in observations before 1840 in Fig. 4 seems suspicious due to the unknown observation hours. The comparison with instrumental data here - while interesting - is not essential, moreover as in any case temperatures before ca. 1850 may contain biases (Melvin et al., 2013; Grudd, 2008).

*(corrected and original observations are now shown in Fig.4)*
Figure 4. Comparison of Torneträsk temperature reconstruction and Tornedalen temperature observations 1802-1977. Reconstruction mean has been adjusted to fit the corresponding observational mean during 1850-1950 and both curves have been smoothed.

(revised version of the discussion, page 7, line 17 – page 9, line 7)

4.1. Bias of the Torneträsk reconstructions

In Rinne et al. 2014 it was observed that both in the nearby oceanic (August SST, Norwegian Sea, Miettinen et al. 2012) and continental (Esper et al. 2012) temperature estimates of the long term oscillations clearly and similarly differ from those derived from the Torneträsk data. Such local anomalies are climatologically impossible and therefore the
mutually similar long-term oscillations in Torneträsk reconstructions contain a bias. Accordingly, those reconstructions are suitable for our error studies.

The differences observed are extreme being of the order of magnitude of the greenhouse warming. The long-term bias in Torneträsk reconstructions is thus detected climatologically. In our computations we estimate that climatological bias as the difference between the Torneträsk reconstruction and the corresponding Esper et al. (2012) reconstruction, the latter having a high number of trees. In Fig. 3a the corresponding anomalies of the Torneträsk reconstruction in Briffa et al. (1992) are compared with our results. The climatologically biased quasi cycle of ca. 350 years is well seen. Especially the recent years are similarly biased showing an underestimation of the temperature. Excluding this recent underestimation, the variants of reconstructions in Melvin et al. (2013, their Fig. 4) show a similar long-term structure. Importantly, the bias is not in the tree ring measurements but in the reconstructions. Hence the bias has essentially remained as such from Briffa et al. (1992) until the present time, in spite of the developments in reconstruction methods. Naturally there can be differences in details because of different computational approaches.

4.2. Explanation of the bias observed

The error formulae derived are of a general form (Eqs. 4 and 5). The error variances depend on the relative error terms and the dimensional portions \((b_{\text{max}} r_i)^2, (b_{\text{max}} r_n)^2\) and \(S^2\). Because the relative error terms depend only on the distribution of the measurements over the age classes and years, they can be applied to any data without knowing the measurements or reconstruction method used.

In Fig. 3 the relative variances are applied for the Torneträsk data with 65 trees. Paucities in the tree population or periods of missing trees (Fig. 1) cause successively data gaps in young, intermediate and old age classes leading to variance variations in Fig. 3d. The increased variance decrease the degrees of freedom and so makes the bias more probable, which is then seen as erroneous long-term oscillations in the reconstructions.

To illustrate the combined impact of different error terms, the degrees of freedom are computed with the aid of the approximate formula in Eq. (7) and are shown in Fig. 3c. In the beginning of the data, they grow slower than the number of measurements (Fig. 3c), show rapid changes thereafter and are often very low. Main minima and maxima of
the DOF are indicated by full and open circles, respectively. Fig. 3b shows the bias estimate of the reconstruction with $b_{\text{max}}=270$. It is seen, that strong minima (maxima) of the DOF in Fig. 3c well predict following biased (unbiased) values in Fig. 3b and fit with the anomalies in reconstructions shown in Fig. 3a.

Low values of the DOF (<5) are related to extrema of the bias. High values of the DOF (>14) indicate vanishing bias. The latter ones are not numerous and therefore bias-free cases are seen only temporarily in Fig. 3b.

One year with low DOF is separately indicated in Fig. 3. A change in the tree population is seen in 1356 AD (Fig. 1) where the oldest age class (<270 years) drops suddenly down to $b=116$ years. As a consequence, variances of the older and intermediate age classes are peaked in Fig. 3d, further lowering strongly the DOF in Fig. 3c. This is reflected in the bias (Fig. 3b) and in the reconstruction (Fig. 3a, smoothed values in Figs. 3a and 3b naturally lag the sudden changes in Figs. 3c and 3d). Correspondingly the optimal distribution of the measurements around 869 and 1500 AD (Fig. 1) are reflected in low error variances (Fig. 3d), high number of the DOF (Fig. 3c) and unbiased temperature estimates (Fig. 3b).

In Fig. 1, there is a repetitive similarity between the two longer periods with nearly continuous samples of trees (AD 1250-1500 and 1600-1800) and two periods without new trees (AD 1500-1600 and 1800-1980). The corresponding terms of the relative variance (Eq. 5) will be discussed in the following.

The large first peaks of variance around 1356 AD (Fig. 3d) result from missing intermediate and old age classes. These damp slowly out and new peaks are seen at 1600 AD. Here they are due to the intermediate (interpolated) and young age classes. The same structure is repeated between 1600-1800 AD and 1800-1980 AD. In this way the paucities due to missing trees in the data cause alternating data gaps leading to varying behavior in temperature reconstructions. Especially biases around 1600 and 1950 can be explained by the error terms connected to the missing younger age classes (due to the missing trees). The latter case is known as “divergence”, a systematic underestimation of the temperatures (Briffa et al. 1998, D’Arrigo et al. 2008). As there is no formal difference between the cases of 1600 and 1950, it is natural to refer to "divergence" in both cases.
4.3 The performance of the reconstruction method applied

The reconstruction method has been designed for estimation of the error variances. A general form of variances is found in Eq. (5). This is applicable as well for tree ring widths as maximum latewood densities. Only three parameters are needed in explicit applications \((r_1, S, r_n)\). Error sources are not corrected but retained in order to study them. For instance, the long-term erroneous oscillations in Briffa 1992 are well reproduced in Fig. 3a during recent years, too. As a summary, the reconstruction method performs as it should.

The reconstruction method gives two values for every year, an average and its error estimate. If the parameters \((r_1, S, r_n)\) were known, yearly confidence limits of the reconstruction could be given. An important part here is the variances due to missing young and old age classes.

To illustrate, the Torneträsk data is applied. Figs. 3a and 4 show that the reconstruction method performs sufficiently satisfactorily and so the error estimates can be seen to be reasonable. On the other hand, the relative errors due to data gaps and paucities are seldom low (e.g. years 869 and 1500 in Fig. 3d). Therefore the reconstruction can be expected to be inaccurate during most years.

If the errors due to the distribution of the measurements are not taken into account, in Eq. (6) \(Var_{rel}=1\). Conventionally the sample accuracy is then characterized by presenting the yearly number of measurements. This practise is followed here. Instead of trying to estimate confidence limits, the sample accuracy is characterized by giving the degrees of freedom (Fig. 3c). The practical application requires that the general error terms in Eq. (5) are approximated and compressed into Eqs. (6) and (7).

In the recent literature new methods to estimate the age dependence function more precisely are introduced (e.g. Briffa et al. 2013; Melvin et al. 2013; Matskovsky and Helama 2014). They improve the reconstruction and the errors in the long-term oscillations may be decreased. Potential error sources in the sampling technique have been detected (e.g. Bowman et al. 2013). As far no new trees are added, the distribution of the measurements is not changed. Hence \(S^2\) in Eq. (6) will be decreased but \(Var_{rel}\) in Eq. (7) is unchanged. The latter term describes the error source studied here and depends only on the
distribution of the measurements. Accordingly, it is more generally valid and is independent of developments in the reconstruction methods and their age dependence functions.

In our reconstruction method the missing younger and older age classes are taken into account. It is natural that the results resemble to those that use age dependence functions. However, here the impact of the missing younger and older age classes is estimated directly from the measurements and only for the average impact of the missing age classes. The resulting estimates of error variances are high in the cases of data gaps and paucities. Therefore there is here no need for a more advanced reconstruction method because necessarily the error estimates will be high and the estimated impact of missing age classes less accurate.

There is one possibility in some specific cases to regulate the accuracy. Our method implies that an upper limit of the age classes is selected. The higher is the limit, the more measurements are included into the analysis. Simultaneously there will be more longer data gaps. It is to find a compromise between the opposite effects. An illustration is given in Fig. 1, where the cases of \( b_{\text{max}} = 270 \) and \( b_{\text{max}} = 370 \) are compared. In the Tornedalen case study, it was concluded to make use of \( b_{\text{max}} = 370 \) in order to decrease the bias during recent years. That choice was motivated as the hatched line in Fig. 3a is closer to zero. Otherwise it is seen that the changes are weak and the selection of \( b_{\text{max}} \) is not decisive. If it wished to be made mechanically, a possibility would be to minimize \( \text{Var}_{\text{ave}} \) with respect to \( b_{\text{max}} \) in Eq. (7). In Esper et al. (2014), the upper limit was taken to be \( b_{\text{max}} = 306 \).