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Response to comment:

“On misleading solar-climate relationship

by B. Legras, O. Mestre, E. Bard and P. Yiou,

Clim. Past Discuss., 6, 767-800, 2010”.

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25 We respond to the comments formulated by Legras et al (2010, hereafter referred to
26 as LMBY), on two papers published in the *Journal of Atmospheric and Solar Terrestrial*
27 *Physics* (Le Mouél et al, 2010a, hereafter referred to as LMKC; and Kossobokov et al,
28 2010, KLMC) on evidence and statistical significance of solar signatures in some of the
29 longest time series of temperature measurements available in Europe. LMBY provide
30 Mathematica codes in their Supplementary material, which can be used as a tool-box by
31 other parties interested in trying to check and reproduce our results and obtain their own
32 independent evaluation of the measurements available, in particular the air temperatures
33 provided, among other data, by the ECA team (Klein Tank et al, 2002;
34 <http://eca.knmi.nl>). The three conclusions of LMBY as summarized in their abstract are:
35 “1) that correlation with solar forcing alone is meaningless unless other forcings are
36 properly accounted and that sunspot counting is a poor indicator of solar irradiance, 2)
37 that long series of temperature require homogenization to remove historical artefacts that
38 affect long term variability, 3) that incorrect application of statistical tests leads to
39 interpret as significant a signal which arises from pure random fluctuations”. We believe
40 all of these conclusions can either be rejected, or are valid but do not apply to our work,
41 as discussed below.

42 LMBY’s introduction contains a series of rather general statements which are not
43 specifically addressed to our papers but which nevertheless deserve brief comments. For
44 instance, the lack of multi-decadal trend in either total solar irradiance or the UV part of
45 the solar spectrum cannot be ascertained from the present (mostly satellite based) data
46 which are available only for the past three to four decades. This is well illustrated by
47 ongoing controversies, such as that between Scafetta and Willson (2009) and Krivova et
48 al. (2009). We also strongly disagree with the statement that “*Correlations are a basis of*
49 *knowledge in areas (...) where the theory is qualitative and does not provide*

50 *mathematical tools for prediction*". We believe that science must rely on a balanced mix
51 of observation, theory and numerical modeling, and that this balance may not be
52 sufficiently maintained in certain areas; also, correlations can be, in almost all disciplines
53 and certainly in geophysics, a very powerful guide to the advancement of understanding
54 and formulating theories, and are part of the observational checks that allow to verify
55 them.

56 Before we start, we would like to quote in full several paragraphs from Press et al
57 (1992) which we try to adhere to as we believe everyone should:

58 *"Data consist of numbers, of course. But these numbers are fed into the computer,*
59 *not produced by it. These numbers to be treated with considerable respect, neither to be*
60 *tampered with, nor subjected to a numerical process whose character you do not*
61 *completely understand. You are well advised to acquire a reverence for data that is*
62 *rather different from the "sporty" attitude that is sometimes allowable, or even*
63 *commendable, in other numerical tasks.*

64 *The analysis of data inevitably involves some trafficking with the field of statistics,*
65 *that gray area which is not quite a branch of mathematics - and just as surely not quite a*
66 *branch of science. In the following sections, you will repeatedly encounter the following*
67 *paradigm:*

68 . *apply some formula to the data to compute "a statistic"*
69 . *compute where the value of that statistic falls in a probability distribution that is*
70 *computed on the basis of some "null hypothesis"*
71 . *if it falls in a very unlikely spot, way out on a tail of the distribution, conclude that*
72 *the null hypothesis is false for your data set.*

73 *If a statistic falls in a reasonable part of the distribution, you must not make the*
74 *mistake of concluding that the null hypothesis is "verified" or "proved". That is the curse*
75 *of statistics, that it can never prove things, only disprove them! At best, you can*
76 *substantiate a hypothesis by ruling out, statistically, a whole long list of competing*
77 *hypotheses, every one that has ever been proposed. After a while your adversaries and*
78 *competitors will give up trying to think of alternative hypotheses, or else they will grow*
79 *old and die, and then your hypothesis will become accepted. Sounds crazy, we know, but*
80 *that's how science works! »*

81 “1) correlation with solar forcing alone is meaningless unless other forcings are
82 properly accounted and that sunspot counting is a poor indicator of solar irradiance”. In
83 our papers we have attempted to test the hypothesis that a solar signature can be
84 recognized in long records of temperatures (or in transforms of these records) from
85 European stations. We have been looking for possible evidence of forcings with, for
86 instance, 11-yr quasi periodicity, which of course means we have not been mixing this
87 with other, non periodical forcing factors. And we have in no way implied that the
88 mechanism was fully understood (existence versus unicity of a solution) but we have
89 made suggestions that could lead to identifying such mechanisms. This certainly does not
90 in itself require full modeling of all forcings, in particular if one keeps in mind that some
91 forcings may yet have to be discovered and properly modeled, and that several depend on
92 key parameters (sensitivities, feedbacks) that are highly uncertain (e.g. Lindzen, 1999).
93 The present view of the majority of scientists involved in climate change is that forcings
94 include solar effects, but with parameters that lead them to conclude that the contribution
95 of such solar forcing is minor. Regarding the use of sunspots as an indicator of solar
96 activity, we are well aware of the fact that the former does not reflect the latter in a one to
97 one correspondence. But, as we have shown for instance in Le Mouël et al (2009, Figure
98 4), the long term (multi-decadal) changes in sunspots, the aa-index and a number of
99 geomagnetic proxies all behave in the same important general pattern with a succession
100 of rising, decreasing, rising and then decreasing again segments over the 20th century
101 (what we had called the “*overall magnetic tendency*” in Le Mouël et al (2005) and which
102 should better be termed the “*secular variation of solar activity*”) (see also Blanter et al,
103 2006; Le Mouël et al, 2007; Shnirman et al, 2009). In that particular case, should formal

104 correlation coefficients be used with so few degrees of freedom, most tests would likely
105 fail (indicating that rigorous statistics is not always the best guide for physical insight).
106 Regarding for instance the effect of volcanic eruptions (which we have some experience
107 with, e.g. Chenet et al, 2005), we note that their effects tend to last only a few years and
108 significantly less than a solar cycle, and also that, whereas some eruptions apparently
109 coincided with periods of low sunspot numbers, three major eruptions occurred in the
110 second half of the 20th century which LMBY note is a period of high (multi-decadal)
111 solar activity.

112 The international sunspot numbers are the only and, therefore, the best available
113 proxy of solar activity in the last 250 years. Better proxies of solar activity, such as solar
114 irradiance, cover at most the past three or four solar cycles and, regretfully, cannot be
115 used in the study we have undertaken (LMKC, KLMC).

116 “2) *long series of temperature require homogenization to remove historical*
117 *artefacts that affect long term variability*”. The question of homogenization of the data
118 has already been raised by Yiou et al (2010) and we have responded to it (Le Mouél et al,
119 2010b). We repeat here that we have a profound disagreement with these authors’ view of
120 what they call data homogenization, based among other reasons, on decades of
121 experience with magnetic observatory data that raised the same questions (baseline
122 change, change in instrument or measurement location or observer’s practice, influence
123 of local magnetic anomalies, numerical errors in transferring the data to a database,...;
124 e.g. Le Mouél et al, 2004; Chulliat et al, 2005). We warn against automated correction
125 and homogenization of these data: it is highly unlikely that one can do better (except of
126 course for easily detected very large errors) than the original observers, particularly for

127 old data. Proper homogenization of data would imply extended stays in the original
128 observatories. Yiou et al (2010) had already noted the existence of homogeneity problems
129 in the data base and pointed out that “more than 94% of stations are flagged as “*doubtful*”
130 or “*suspect*””. The definition of data quality and suspect stations should be specified: the
131 ECA&D data base has three quality control checks for data values: “*Flag=0*” or “*valid*”,
132 “*Flag=1*” or “*suspect*” and “*Flag=9*” or “*missing*”. Stations are put in three classes
133 according to four homogeneity tests: “*Class 1*” or “*useful*” when no more than 1 test
134 rejects the null hypothesis at the 1% level, “*Class 2*” or “*doubtful*” when 2 tests reject the
135 null hypothesis at the 1% level and “*Class 3*” or “*suspect*” when 3 or 4 of the four tests
136 reject the null hypothesis at the 1% level. Note that due to the definition used by the
137 database editors, the best data in their data base are designated as “*suspect*”! LMBY share
138 these views and recommend performing a data massage such as “*homogenization*” in
139 order to smooth out the roughness of unusual details. We are against this practice. Table 1
140 in LMBY provide a special tabloid “*excerpt*” on homogeneity checking results for the
141 1901–2007 period pointing to the five “*suspect*” stations, but not reminding readers that
142 only 7 out of 126 stations in Europe are found to be “*useful*”, whereas 118 are “*suspect*”
143 and 1 “*doubtful*”! How could many of the results we and others have found based on
144 these European observations have been obtained with suspect data containing only noise
145 and no useful information? And how many useful stations are there in the world? It is
146 clear that the homogeneity checking, which rejects 95% of all European data and stations,
147 as recognized by LMBY, is not a useful or reasonable test. LMBY also point out a
148 suspect potential artefact in Bologna, which we had also noted and discussed. LMBY
149 note that there was a change of thermometer in 1867 and a relocation “*to a different*

150 *place*” (not mentioned by the ECA team) in 1881. These “*meta data*” do not fit the
151 “*artefact*”, which begins in 1865 and ends before 1880: the dramatic changes in
152 (maximum) TX, i.e. rise and fall, were recorded using the same thermometer and at the
153 same place, and thermometers did not show any suspect behavior in variation of
154 (minimum) TN. Finally, we repeat that “*TN and TX values are all of the highest quality*
155 *code in ECA at each of the three locations*”, i.e., “*Flag=0*” that is “*valid*”. As a
156 conclusion, we reject LMBY’s assertion that “*detection and correction of these*
157 *heterogeneities are absolutely necessary before any climate study can be based on the*
158 *instrumental series*”. Of course it is better to correct obviously erroneous data points, but
159 we reject blind systematic corrections. What our studies show is that using the raw data
160 does produce interesting information that cannot be due to noise or chance.

161 “3) *incorrect application of statistical tests leads to interpret as significant a signal*
162 *which arises from pure random fluctuations*”. This general “*lesson*” is of course right, but
163 cannot be implied as a criticism of our studies. On the contrary, we have identified a
164 major error in LMBY which invalidates most of their remaining criticism on statistical
165 validity of our results.

166 Some of the main issues of statistical significance are left for the second part of
167 KLMC, which LMBY have left for future analysis elsewhere. Without waiting for this
168 new piece of comment, we wish to simply refute here our commentators’ claim that we
169 have been using the *standard error* in an erroneous way in order to ascertain the
170 separation of the averages. LMBY have been splitting our averages into “*averages of*
171 *averages*” and have thereby created confusion in quantifying the statistical significance of
172 potential solar signals. In fact, the sizes of samples $\{T_H ik\}$ and $\{T_L ik\}$ are $21 \times N_H$ and

173 $21 \times N_L$, and the standard errors on these involve the square of these quantities in the
174 denominator. LMBY have erroneously reduced these numbers to N_H and N_L , i.e. the
175 numbers of years in the periods H and L , which leads them to use $21 \times N_H^2$ and $21 \times N_L^2$ in
176 estimating the standard error of the mean of the mean, $E(E(T_{ik}))$. LMBY therefore
177 overestimate confidence intervals by a factor of about square root of 21.

178 In addition, as can be seen in their supplementary material, when trying to account
179 for dependencies in a 21-day interval (which we select), LMBY use 90- and 150-day
180 intervals that naturally are affected by the seasonal variability of temperatures (plot and
181 output on page 21, SM to LMBY). Figure 1 actually shows that autocorrelations of the
182 daily temperatures in 21-day intervals fall below 0.2 in less than 3 days, while
183 autocorrelation for the daily range of temperatures ΔT (which LMBY fail to consider)
184 falls below 0.2 on the second day. This is why we used a 21-day interval and not a longer
185 one biased by seasonal effects (from LMKC: “*A 21-day (i.e. 3-week) centered moving*
186 *average is applied: indeed, this is both long enough to stabilize the still noisy averaged*
187 *calendar values and yet short enough that features with monthly and longer time*
188 *constants are well preserved.*”). The effective number of degrees of freedom in our
189 estimates is therefore not reduced as thought by LMBY. As a consequence of their error,
190 the attempts of LMBY to use the T-test technique are biased by the LMBY 9-day
191 dependencies attributed to weekly variability of air temperatures (LMBY page 780 and
192 supplementary material). Should this 9-day correlation apply to our Planet Earth, in our
193 Solar System, weather forecasts would be greatly improved... The sentence “*the number*
194 *of effective degree of freedom is about 9 times smaller than estimated by LKMC and*
195 *consequently the estimated variance of the ensemble average is about three times larger*”

196 is therefore false.

197 The checks on our two papers in response to LMBY lead us to conclude that their

198 criticisms are either irrelevant or erroneous and to reaffirm all of our conclusions.

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202 Acknowledgements: IPGP Contribution NS xxxx.

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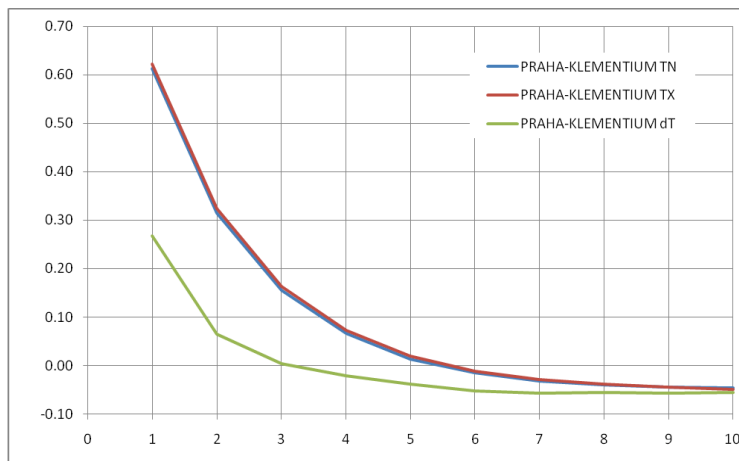
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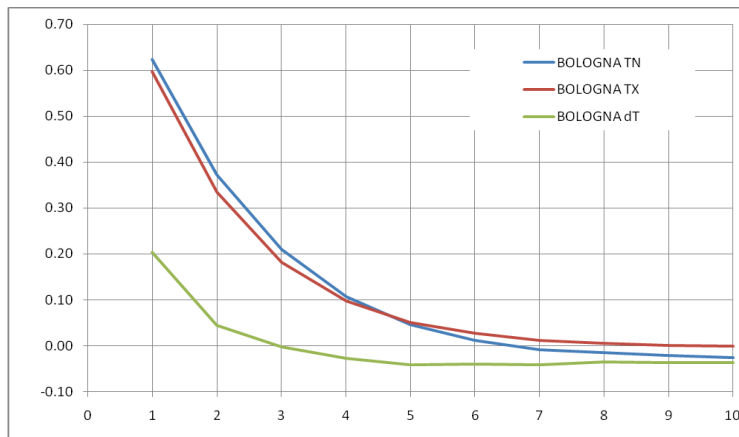
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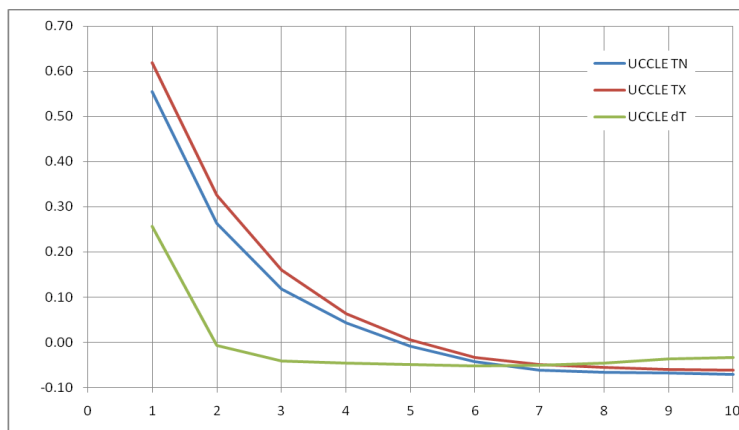


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Figure 1.
Autocorrelation functions of TN, TX and $\Delta T = TX - TN$ at Prague, Bologna, and Uccle.